

# Quantitative Methods in Political Science: Sampling and Statistical Inference

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## Roadmap

- Understand and model stochastic processes
- **Understand statistical inference**
- Implement it mathematically and learn how to estimate it
  - OLS
  - Maximum Likelihood
- Implement it using software
  - R
  - Basic programming skills

## Statistical inference

- Basics

- Sampling

- Example

## Confidence Intervals

- Construction of CIs: Normal approximation

- Construction of CIs: Bootstrapping

- Construction of CIs: Simulation

## Testing Hypotheses

# Statistical inference

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# Statistical Inference: the Example of Polls

The screenshot shows the Politics section of The New York Times website. At the top, there are navigation links for HOME PAGE, TODAY'S PAPER, VIDEO, MOST POPULAR, and TIMES TOPICS. The main header includes the newspaper's name, date (Tuesday, September 18, 2012), and the word "Politics". A search bar is located in the top right corner. Below the header is a horizontal menu with categories: WORLD, U.S., N.Y./REGION, BUSINESS, TECHNOLOGY, SCIENCE, HEALTH, SPORTS, OPINION, ARTS, STYLE, TRAVEL, JOBS, and REAL ESTATE. A row of icons represents various political topics: POLITICS HOME, THE CAUCUS, FIFTYTHIRTY EIGHT, ELECTION 2012, G.O.P. PRIMARY, INSIDE CONGRESS, and POLL WATCH. A large advertisement banner for Somerset by Alice Temperley and John Lewis is displayed. Below the banner is a section titled "The Caucus" with the subtitle "The Politics and Government Blog of The Times". The article "Poll: Obama Holds Narrow Edge Over Romney" by Jeff Zeleny is featured, dated September 14, 2012, with 661 comments. The article text states: "President Obama holds a narrow three-point advantage over Mitt Romney among Americans most likely to vote in November, according to the latest New York Times/CBS News poll." To the right of the article is a "Search This Blog" box, a "Previous Post" link to "The Early Word: In It for the Duration" by Obam Ceremon America, and social media links for Facebook and Twitter. A "THE AGENDA" section at the bottom right mentions that Times reporters and editors examine big issues facing the country.

“...Including those who lean toward a specific candidate, the president has 49 percent and Mr. Romney has 46 percent, a difference within the **margin of sampling error** of plus or minus three percentage points on each candidate...” *New York Times*, Sept 18, 2012.

# The Goal of Statistical Inference

- Learn a quantity of interest about a particular group (**population parameters**).
  - Income of working age population in a country.
  - Reading ability of 4th graders in France.
  - Support for environmental policy in Germany.
- Often information on all members of the group (**population**) will not be available.
- We use **sampling** to collect a limited amount of information and use it to **infer** population properties (**parameters**).
- Since we have only information on a **subset** of the population we are **uncertain** about our inference (but there are other sources of uncertainty as well even if we observe the entire population).
- All inferences are inherently uncertain!

## Statistical Inference

The goal of statistical inference is to estimate **population parameters** and summarize our uncertainty about these estimates.

# Terminology

- A **parameter** describes a feature of the population. The parameter is **fixed** at some value, and we will **never** be able to know it for sure.
- What we observe is a **random sample**, drawn from the population. A random sample is a proper subset of the population for which it is true that each member has an **equal probability** to be selected.
- From this sample, we can calculate a **sample statistic** of a population parameter. A sample statistic is a function that is applied on the observed sample. This function is called the **estimator** of the population parameter.
  - We can calculate the mean of a random sample. The mean is then a *sample statistic*, and the function that maps observations of the random sample to this sample statistic is the *estimator*.
- The population parameter is the **estimand**. The result of such a calculation is called an **estimate**.
  - Note: The estimator is a function, the estimate is a number!

# The Principle of Sampling

- Probability sampling: Select from a population with size  $N$  a number of individuals,  $n$  (usually  $n \ll N$ ), such that each individual has a non-zero probability of being chosen.
- Sources of variation across samples:
  - **Sampling variability**: Means and standard deviations of repeated samples will not be identical.
  - **Sampling error**: An estimate from a sample will not be identical to the value in the population.
- The sample size is positively related to the desired precision of the estimate.



## Sampling model:

- We have: (1) a population, (2) a sample from this population, and (3) an estimate of a population parameter.
- How uncertain are we about that estimate?
- Alternatively, how precisely can we estimate the population parameter?
- In a different sample, our estimate would be slightly different. Hence, estimates vary over **repeated samples**.
- Applying an estimator on repeated samples yields a **sampling distribution** for this statistic.
- Calculating the spread of this sampling distribution yields a **measure of uncertainty**.

## Example of Statistical Inference

- Let there be a country with 100,000 inhabitants.
- We want to know what the mean income of this country is.
- We sample 5000 individuals randomly from the population.
- The mean of our obtained sample is 1400 with a standard deviation of 2000.
- The standard error of the estimate of the population mean ( $\hat{\theta}$ ) is:  $\sigma_{pop}/\sqrt{n}$ .
- For a **large sample**, the standard deviation ( $\hat{\sigma}$ ) of a sample can be used as an **approximation** of the population standard deviation ( $\sigma_{pop}$ ):

$$SE(\hat{\theta}) = \frac{\hat{\sigma}}{\sqrt{n}} = \frac{\widehat{2000}}{\sqrt{5000}} = 28.3$$

- The mean income of our population is  $1400 \pm 28.3$

## Digression: Derivation of Standard Error

- Let's assume that we have a **random sample** from a population, i.e., we have  $n$  random variables,  $\theta_1, \dots, \theta_n$  that come from the same population represented by a distribution with mean  $\mu$  and variance  $\sigma^2$
- Such random variables are called *independently and identically distributed (iid)*
- Hence, we know that  $\text{Var}(\theta_i) = \sigma^2$  for all random variables  $\theta_i$ .
- Denote their mean (i.e., sample average) as  $\bar{\theta}$ .
- Then, we have

$$\begin{aligned}\text{Var}(\bar{\theta}) &= \text{Var}\left(\frac{1}{n} \sum_{i=1}^n \theta_i\right) = \frac{1}{n^2} \text{Var}\left(\sum_{i=1}^n \theta_i\right) = \frac{1}{n^2} \sum_{i=1}^n \text{Var}(\theta_i) \\ &= \frac{1}{n^2} n \sigma^2 = \frac{1}{n} \sigma^2\end{aligned}$$

- Hence, the standard error of the mean is derived as  $\text{SE}(\bar{\theta}) = \frac{\sigma}{\sqrt{n}}$
- The sampling variance equals the population variance divided by the sample size.

# Confidence Intervals

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# Confidence Intervals

- Our estimate of a population feature (parameter) varies across repeated samples, thus generating a *sampling distribution*.
- Instead of a **point estimate**, we should better get an **interval estimate** – a range within which the true parameter lies with some level of certainty.
- Using the standard error or the variance of our estimates we can construct **confidence intervals**.
- We call a confidence interval a  $q\%$  confidence interval if it is constructed such that it contains the true parameter at least  $q\%$  of the time *if we repeat the experiment a large number of times*.
- A 95% confidence interval hence contains the true parameter at least 95% of the times *if we repeat the experiment a large number of times*.
- Note that this **does not** mean that there is a 95% probability for the population parameter to lie inside the interval!

# Three Different Approaches to Assess Uncertainty

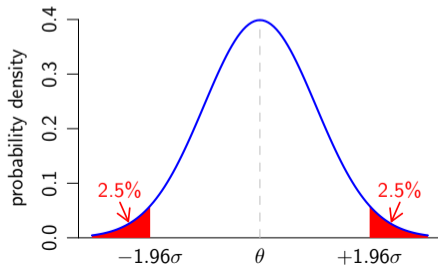
1. Analytical
2. Bootstrapping (resampling)
3. Simulation (parametric)

# 1. Analytical

# Analytical Approach: CIs via Normal Approximation

- We have a sample statistic  $\hat{\theta}$  estimated for a parameter  $\theta$ .
- If the **sample is large enough**, we can assume a normal sampling distribution with mean  $\theta$  and variance  $\text{Var}(\hat{\theta})$
- We can then construct a 95% confidence interval using the quantiles from the standard normal distribution:

$$\hat{\theta} \pm 1.96\sqrt{\text{Var}(\hat{\theta})} = \hat{\theta} \pm 1.96 \cdot \frac{\hat{\sigma}}{\sqrt{N}}$$





## Example: Measurement With Normal Error

- Support for a new policy during the last month as measured by a polling firm ( $N = 50$ ), as

15 16 12 17 14 13 15 16 12 14 17 15 12 15 14 16 16 14 13 12 13 15 16 14 15

11 13 13 16 15 17 14 12 15 14 13 16 17 14 15 16 14 13 14 13 15 17 11 14 15

with an average value of  $\hat{\theta} = 14.36$  and a standard deviation of  $\hat{\sigma} = 1.61$ .

- The **standard error** of the estimate  $\hat{\theta}$  of the population mean is

$$SE(\hat{\theta}) = \frac{\hat{\sigma}}{\sqrt{N}} = \frac{1.61}{\sqrt{50}} = 0.228$$

- The 95% confidence interval ranges from

$$14.36 \pm 1.96 \times 0.228 = [13.91, 14.81]$$

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## 2. Bootstrapping (resampling)

# Bootstrapping Approach: Constructing CIs via Resampling

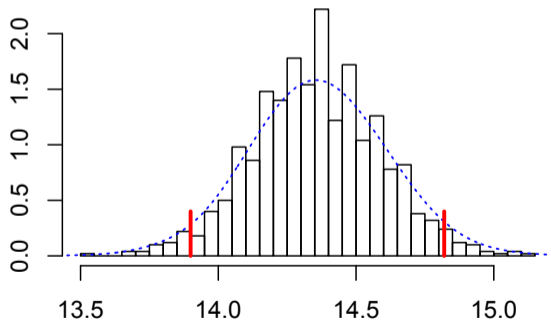
- What if the population does not appear normal and/or we have only a small sample?
- We do **not know** what the sampling distribution looks like, but we want a (robust) estimate of it.
- **Bootstrapping** estimates the sampling distribution of  $\theta$  by **repeatedly sampling** (*with replacement*) from the original sample.
- In standard bootstrapping, the size of the bootstrap sample,  $n$ , is **identical** with the sample from the population. The variability comes from **sampling with replacement**.
  1. Take  $s$  samples of size  $n$  from your data.
  2. Calculate the quantity of interest ( $\hat{\theta}_i$ , e.g. the mean) for each of your  $s$  samples, which yields a vector of length  $s$ .
  3. A simple confidence interval for your quantity can be obtained by calculating quantiles (e.g., 2.5 and 97.5 percentiles for 95% CI) of this vector.

## Just remember!

The population is to the sample as the sample is to the bootstrap sample.

## Example cont.: Measurement with Normal Error Revisited

- Bootstrapping  $s = 1000$  samples of size  $n = 50$  yields a mean of 14.36256.
- A simple confidence interval can be obtained by just calculating the 2.5th and 97.5th quantile of the bootstrapped sampling distribution.
- This yields a 95% confidence interval from 13.8995 to 14.8200.



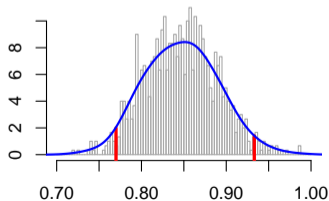
### 3. Simulation (parametric)

# Simulation Approach: Constructing CIs via Simulation

- An alternative **parametric approach** sets up a population distribution and draws from it.
- This is an empirical / computational implementation of the idea of inference as **repeated sampling**.
- We obtain **hypothetical** repeated samples from a population distribution.
- These **simulations** are done using a computer:
  1. Create a (normal) sampling distribution from the mean and standard error of your sample.
  2. Take  $s$  draws from that distribution  $N(\hat{\theta}, \hat{\sigma}^2)$ .
  3. Calculate your quantity of interest  $s$  times. Thus, we simulated its **sampling distribution**.
  4. Calculate summaries, such as means and standard errors, for the resulting vector of length  $s$ .

## Example: Proportions

- Consider a sample of  $n = 1000$  individuals, 500 of which are men and 500 are women.
- 55% of men vote left, while 65% of women vote left.
- *Quantity of interest*: The ratio for voting left for men compared to women is 0.846 ( $\approx \frac{0.55}{0.65}$ ).
- Analytically, the SE for proportions is  $SE(\hat{p}) = \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$ .
- The standard error of these proportions for men is 0.0222 and for women is 0.0213.
- The confidence interval for this ratio obtained via simulation ranges from 0.765 to 0.937.



# Example: Proportions

## Mechanics:

- Take  $s = 1000$  draws from normal distribution with  $\hat{\mu}$  and  $\hat{\sigma}^2$  like the data

$$p_m = \mathcal{N}(.55, .0222^2) = [0.56804, 0.53736, 0.53876, 0.55186, 0.55523, \dots]$$

$$p_w = \mathcal{N}(.65, .0213^2) = [0.63091, 0.63087, 0.61234, 0.64373, 0.68099, \dots]$$

- Calculate vector of 1000 ratios between men and women

$$r = p_m/p_w = [0.90036, 0.85178, 0.87984, 0.85728, 0.81533, \dots]$$

- The empirical confidence interval is the 2.5% and 97.5% quantile of the distribution of this vector of ratios (and ranges from 0.765 to 0.937).



# Testing Hypotheses

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# Hypothesis Testing

- Usually, we want to learn about population parameters based on sample statistics. This requires us to use **statistical inference**.
- For example, is a population parameter different from zero? If you find a value different from zero in your **sample**, can you be sure that you did not select a sample where this is the case while in the **population** the value is zero?
- Formulate competing hypotheses:  
 $H_0$ : The population parameter is equal to zero (**null hypothesis**).  
 $H_A$ : The population parameter differs from zero (**alternative hypothesis**).
- Make a decision based on
  - (a) Confidence intervals.
  - (b) Test statistic.

## (a) Decision based on Confidence Intervals

- Someone claims the mean income of a country is 1400. You take a random sample of size  $n=1000$  and obtain a mean income of 1350, with a standard deviation of 750.
- What can we say about the “truth” of this claim?
- Construct two competing hypotheses:  
 $H_0 : \mu = 1400$  and  
 $H_A : \mu \neq 1400$
- Calculate the 95% confidence interval for the mean:

$$1350 \pm 1.96 \times \frac{\widehat{750}}{\sqrt{1000}} = [1326, 1396]$$

- The  $H_0$  value of 1400 does not lie in this interval: the sample is **very unlikely** to come from a population with mean 1400.
- This allows us to reject  $H_0$ .
- Note that we can **always only reject or fail to reject** hypotheses, but we can **never prove** hypotheses.

## (b) Decision based on a Test Statistic

- At a level of 95% confidence, the critical z values on a standard normal distribution are  $\pm 1.96$ .
- How much is the sample mean (i.e., 1350) away from the hypothesized population mean (i.e., 1400) in repeated samples? Calculate the z-score of your realized sample:

$$z = \frac{\bar{X} - \mu_{pop}}{\sigma_{pop}/\sqrt{n}} = \frac{1350 - 1400}{750/\sqrt{1000}} = -2.1$$

- This is smaller than the critical value and you would reject  $H_0$ .

